

# The angular correlation function of the *ROSAT* All-Sky Survey Bright Source Catalogue

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## ABSTRACT

We have derived the angular correlation function of a sample of 2096 sources detected in the *ROSAT* All-Sky Survey (RASS) Bright Source Catalogue, in order to investigate the clustering properties of active galactic nuclei (AGN) in the local Universe. Our sample is constructed by rejecting all known stars, as well as extended X-ray sources. Areas with  $|b| < 30^\circ$  and declination  $\delta < -30^\circ$  are also rejected owing to the high or uncertain neutral hydrogen absorption. Cross-correlation of our sample with the Hamburg/RASS optical identification catalogue suggests that the vast majority of our sources are indeed AGN. A  $4.1\sigma$  correlation signal between  $0^\circ$  and  $8^\circ$  was detected with  $w(\theta < 8^\circ) = 2.5 \pm 0.6 \times 10^{-2}$ . Assuming a two-point correlation function of the form  $w(\theta) = (\theta/\theta_0)^{-0.8}$ , we find  $\theta_0 = 0.062$ . Deprojection on three dimensions, using Limber's equation, yields a spatial correlation length of  $r_0 \approx 6.0 \pm 1.6 h^{-1}$  Mpc. This is consistent with the AGN clustering results derived at higher redshifts in optical surveys and suggests a comoving model for the clustering evolution.

**Key words:** surveys – galaxies: active – quasars: general – large-scale structure of Universe – X-rays: general.

## 1 INTRODUCTION

Active galactic nuclei (AGN) can be detected at high redshift owing to their high luminosities and hence they can be used to place stringent constraints on the evolution of large-scale structure over a wide redshift range (see Hartwick & Schade 1989). Our knowledge on the AGN clustering properties comes mainly from large optical UV excess (UVX) surveys for quasi-stellar objects (QSOs) (Boyle, Shanks & Peterson 1988; Mo & Fang 1993; Shanks & Boyle 1994; Croom & Shanks 1996; La Franca, Andreani & Cristiani 1998). Croom & Shanks (1996) analysed the clustering properties of the Large Bright Quasar Survey (LBQS) and combined the results with that obtained from other QSO surveys including the Durham/Anglo-Australian Telescope (AAT) UVX sample. They derived a clustering length of  $r_0 = 5.4 \pm 1.1 h^{-1}$  Mpc at a mean redshift of 1.27. La Franca, Andreani & Cristiani (1998) derived comparable results ( $r_0 = 6.2 \pm 1.6 h^{-1}$  Mpc) by investigating a sample of  $\sim 700$  quasars in the redshift range  $0.3 < z \leq 3.2$ . Comparison of these clustering results in different redshifts rather favours a comoving model for the evolution of clustering. According to this model the amplitude of the correlation function remains fixed with redshift in comoving coordinates as the galaxy pair expands together with the background mass distribution. In contrast, in the stable model

for clustering evolution, AGN trace clumps of mass that have gravitationally collapsed in bound units and have therefore ceased to take part in the general expansion of the Universe. However, most of the AGN samples used in the analyses above, come from pencil beam surveys and contain a large fraction of high-redshift AGN. Georgantopoulos & Shanks (1994) studied the correlation function of a sample of about 200 *IRAS*-selected Seyfert galaxies. They detect a  $2\sigma$  clustering signal on scales less than  $10 h^{-1}$  Mpc. Although their statistics are limited, their results are more consistent with a comoving model. It is evident that there is a pressing need for large AGN samples in the local Universe in order to place tight constraints on the clustering evolution in a broad redshift range.

Analogous clustering studies in X-rays have been scarce but they are potentially interesting as they provide the opportunity to derive information on how the X-ray selected AGN trace the underlying mass distribution. The first direct study (using redshift information) of the correlation function and clustering properties of X-ray selected AGN is that of Boyle & Mo (1993). They studied the local  $z < 0.2$  AGN clustering properties using a sample from the *Einstein* Extended Medium Sensitivity Survey (EMSS). They did not detect clustering at a statistically significant level ( $\approx 1.5\sigma$ ). Vikhlinin & Forman (1995) analysed the angular clustering properties of a set of deep *ROSAT* observations. Their angular correlation length  $\theta_0$  translates, using Limber's equation (Peebles 1980), to a spatial correlation length  $r_0$  which is

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consistent with that derived in optical surveys. Finally, Carrera et al. (1998) using a sample of 235 AGN from two different soft X-ray surveys, the *ROSAT* Deep Survey (Georgantopoulos et al. 1996) and the *RIXOS* survey (Mason et al. 2000) derived the first direct evidence (i.e. in three dimensions) for clustering in an X-ray selected QSO sample. They derived a low spatial correlation length between 1.5 and  $5.5 h^{-1}$  Mpc depending on the adopted model of clustering evolution.

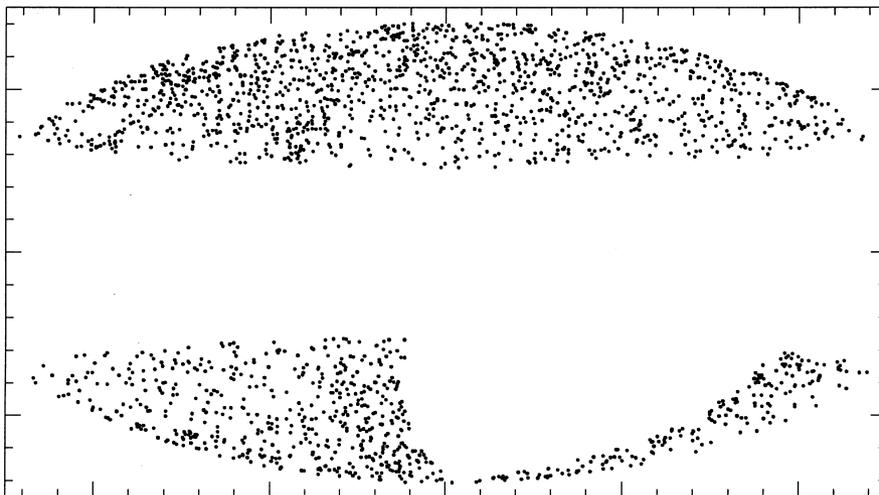
In this paper we present a study of the clustering properties of soft X-ray selected AGN from the *ROSAT* All Sky Survey (RASS). The large number of AGN in this sample provides excellent statistics and hence the opportunity to put stringent constraints on the clustering properties in the local X-ray Universe. Moreover, comparison with clustering results at higher redshift in both the X-ray and optical regime could provide valuable information on the evolution of clustering in the Universe.

## 2 THE *ROSAT* SAMPLE

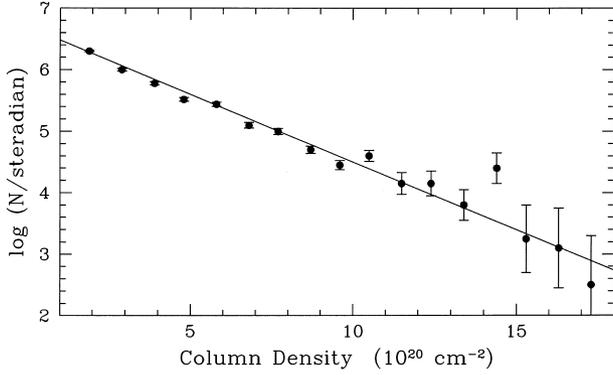
The data used come from the *ROSAT* All-Sky Survey Bright Source Catalogue (hereafter RASSBSC, Voges et al. 1999). The RASSBSC is a subsample of the brightest 18 811 sources in the RASS. The RASS covers 98 per cent of the sky in the 0.1–2.4 keV energy range with 30-arcsec angular resolution. We excluded all sources with count rate less than  $0.1 \text{ counts s}^{-1}$ . This yields a catalogue that has a *uniform* count rate limit over the sky (Voges et al. 1999). We further exclude all sources between the strip  $(-30, 30)$  degrees in Galactic latitude in order to minimize the effect of hydrogen absorption. Although the majority of the RASS sources are AGN there is an appreciable contamination from stars and clusters of galaxies. Therefore in order to study the AGN cross-correlation function we first need to exclude all stars, groups and clusters of galaxies. In principle, one can utilize the Hamburg Optical Identification Catalogue (Bade et al. 1998), which is an ongoing spectroscopic follow-up program of the RASS sources in the northern sky, containing objective prism spectra for several thousand sources. However, as the Hamburg catalogue contains optical identifications only for a small subsample of the RASSBSC, we choose to use a different approach to exclude the stellar and extended objects from our sample. We have therefore cross correlated the RASSBSC with known stellar

catalogues, such as the Smithsonian Astrophysical Observatory Stellar Catalogue (<http://xena.harvard.edu/software/sao>) and the STScI Guide Star Catalogue (<http://www-gsss.stsci.edu>). In addition we have excluded all known *ROSAT* White Dwarfs, Cataclysmic Variables, X-ray Binaries and hot (OB) stars. Details of these *ROSAT* catalogues are given in the GSFC HEASARC web page <http://heasarc.gsfc.nasa.gov>. The above cross-correlation resulted in the rejection of 6246 stars. The next step is to remove all known clusters and groups of galaxies from our sample. Inspection of the Hamburg catalogue shows that all identified clusters appear to be spatially extended in X-rays. Hence we further exclude all sources which have an extension flag value greater than 40 (see Voges et al. for the definition of the extension flag value). Indeed the cross correlation of our sample with the Hamburg RASS catalogue in the common ( $\approx 8000$  degrees) area shows no remaining clusters. After the application of the above corrections the vast majority of our sample, at least in the areas of the sky covered by the Hamburg catalogue, consists of AGN (86 per cent), while there is still some small stellar and galaxy contamination (about 10 and 4 per cent respectively).

Finally, we need to make corrections for the neutral hydrogen in the Galaxy. The 0.1–2.4 keV band is very sensitive to photoelectric absorption. Therefore the apparent sky density may change in different areas on the sky owing to the changes in the hydrogen column density ( $N_H$ ) in our Galaxy. The large-scale variations of the  $N_H$  with Galactic latitude as well as the patchy absorption on smaller scales may introduce spurious signals in the correlation function. Therefore, we used the Leiden/Dwingeloo Atlas of Galactic Neutral Hydrogen, (Hartmann & Burton 1997) in order to remove the effects of the hydrogen absorption. As the above atlas contains  $N_H$  data only for declinations  $\delta > -30^\circ$  we further exclude RASSBSC sources outside the missing regions. The resulting final catalogue contains 2096 sources over 4.9 sr. In Fig. 1 we present the Aitoff projection of our sample in Galactic coordinates. The apparent large-scale density gradients, seen in Fig. 1, are owing to patchy Galactic absorption. In Fig. 2 we present the sky density as a function of the Galactic column density  $N_H$ . The solid line represents the best fit to the data. This is given by  $\log_e N = \alpha + \beta N_H$  where  $N$  is the density of galaxies per steradian and  $N_H$  is the column density in units of  $10^{20} \text{ cm}^{-2}$ . We find  $\alpha = 6.7$  and  $\beta = -0.2$ .



**Figure 1.** Aitoff projection of our final catalogue (2096 sources) in Galactic coordinates. The centre corresponds to Galactic longitude and latitude  $l = 0$ ,  $b = 0$  while  $l$  increases to the left.



**Figure 2.** The AGN surface density as a function of hydrogen absorption  $N_H$ . The error bars represent Poisson uncertainties.

### 3 METHOD AND ANALYSIS

We used the two-point angular correlation function to measure the clustering properties of our sample. The correlation function  $w(\theta)$  is defined in terms of the probability  $\delta P(\theta)$  of finding two objects at a separation  $\theta$  in two small solid angles  $\delta\Omega_1$  and  $\delta\Omega_2$ .

$$\delta P = n^2[1 + w(\theta)]\delta\Omega_1\delta\Omega_2 \quad (1)$$

where  $n$  is the mean number density of this objects. For an unclustered population  $w(\theta) = 0$ , while positive or negative values of  $w(\theta)$  indicate clustering or anticlustering respectively. We estimated the correlation function by comparing the distribution of our sample to that of a random sample. The random sample was constructed by generating angular positions at random over the sky and then folding them through the  $N_H$  map. In particular, the points on the sky with the lowest  $N_H$  are assigned a probability 1 (always accepted) while points on the sky with higher  $N_H$  are assigned lower probabilities, following the  $N$  versus  $N_H$  fit of Fig. 2, and therefore have a lower chance of entering the random catalogue. The correlation estimator factor that we use is

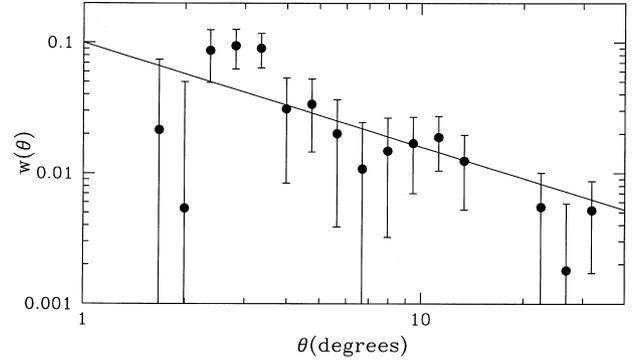
$$w(\theta) = f(N_{dd}/N_{dr}) - 1 \quad (2)$$

while the error is given by (Peebles 1973)

$$\sigma_w = \sqrt{f(1 + w(\theta))/N_{dr}} \quad (3)$$

Here  $N_{dd}$  is the number of data–data pairs and  $N_{dr}$  is the number of data–random pairs for a given separation. We use a random sample 30 times larger than our observed catalogue. Then the normalization factor is  $f = 2N_d N_r / (N_d(N_d - 1))$  the ratio of the total number of independent data–data pairs to the number of data–random pairs. In the equation above  $N_d$  and  $N_r$  are the number of data and random (after folding through the  $N_H$  distribution) points respectively.

The derived correlation function is presented in Fig. 3. We find  $N_{dd} = 26669$  and  $N_{dr} = 26004$  hence detecting a significant clustering at the  $4.1\sigma$  confidence level (using Poissonian statistics) between 0 and 8 degrees with  $w(\theta < 8^\circ) = 2.5 \pm 0.6 \times 10^{-2}$ . On larger scales the correlation function is consistent with an unclustered population. Assuming a two-point angular correlation function of the form  $w(\theta) = (\theta/\theta_0)^{1-\gamma}$  we find that  $\gamma - 1 = 0.9 \pm 0.15$  and  $\theta_0 = 0^\circ.08 \pm 0^\circ.03$ . If the value of  $\gamma - 1$  is fixed at 0.8, which is the value found for optically selected quasars, we find  $\theta_0 = (0^\circ.06 \pm 0^\circ.02)$ . Next, we performed a test in order to check whether our signal is diluted owing to the small



**Figure 3.** The two-point angular correlation function  $w(\theta)$  of our sample (2096 sources).

percentage of stellar contamination. We estimated the correlation function for all the identified AGN in the Hamburg/RASS Catalogue. There were 662 AGN with declination above  $-30^\circ$ . The estimated correlation function is  $w_{AGN}(\theta) = 3.7 \pm 1.0 \times 10^{-2}$ . This is consistent with the  $w(\theta)$  of our sample, within  $\sim 1.2\sigma$ , suggesting that the stellar contamination does not have a major effect on the derived correlation amplitude. We also note that our results do not suffer from the possible ‘amplification bias’. Vikhlinin & Forman (1995) pointed out that when the best-fitting angular correlation length is smaller than the FWHM of the *ROSAT* Position Sensitive Proportional Counter (PSPC) point spread function (PSF), two or more sources can be detected as a single source. As a result the distribution of the confused sources may be biased with respect to the correlation function of the parent population. This ‘amplification bias’ effect works towards the amplification of the correlation function amplitude (cf. Kaiser 1984). However, in our case, confusion problems are very unlikely as the FWHM of the *ROSAT* PSPC ( $\sim 30$  arcsec) corresponds to a very small spatial separation ( $< 100$  kpc) at the typical redshift of the RASS AGN ( $z \sim 0.1-0.2$ ).

Next, we use Limber’s equation which gives the relation between the spatial and angular correlation function in order to calculate the correlation length  $r_0$  in three dimensions. We have used a power-law shape for the spatial two-point correlation function following Groth & Peebles (1977):

$$\xi(r, z) = (r/r_0)^{-\gamma}(1+z)^{-p}, \quad (4)$$

where  $r_0$  is the correlation length in comoving coordinates, the parameter  $\gamma$  is fixed at the value of 1.8 and  $p$  is the parameter describing clustering evolution. If  $p = 0$  the clustering is constant in comoving coordinates (comoving clustering). If  $p = 1.2$  then the clustering is constant in proper coordinates (stable clustering). The amplitude  $\theta_0$  in two dimensions is related to the correlation length  $r_0$  in three dimensions through the equation (Peebles 1980)

$$\theta_0^{\gamma-1} = r_0^\gamma H_\gamma \left( \frac{H_0}{c} \right)^\gamma \frac{\int_0^\infty dy \phi(y)^2 [1 + z(y)]^{-p} y^{5-\gamma} / F(y)}{\left[ \int_0^\infty y^2 dy \phi(y) / F(y) \right]^2} \quad (5)$$

where  $H_\gamma = \Gamma(\frac{1}{2})\Gamma(\frac{\gamma-1}{2})/\Gamma(\frac{\gamma}{2})$ . The parameter  $y$  is related to the distance and redshift through

$$y = 2 \frac{(\Omega - 2)(1 + \Omega z)^{1/2} + 2 - \Omega + \Omega z}{\Omega^2(1 + z)} \quad (6)$$

Finally,  $F(y) = [1 - y^2(\Omega - 1)]^{1/2}$ , while  $\phi(y)$  determines the fraction of sources observable at a given  $y$  (or  $z$ ), i.e., those with

observed fluxes greater than the detection threshold. The cosmological parameter  $\Omega$  enters through the  $y - z$  relation, the volume factor  $F(y)$  and the selection function  $\phi(y)$ ; here we use  $\Omega = 1$ . The selection function  $\phi(y)$  depends also on the evolution of the AGN luminosity function. Boyle et al. (1993) derived the AGN cosmological evolution in the form of pure luminosity evolution, with the characteristic luminosity  $L_x^*$  varying with redshift as  $(1+z)^3$ . The selection function can be written as

$$\phi(y) = \int_{L_{min}}^{\infty} \Phi(L_x, z) dL \quad (7)$$

where the local X-ray luminosity function is given by a double power-law function (Boyle et al. 1993).

The lower limit of integration  $L_{min}$  corresponds to the luminosity that can be observed at redshift  $z$  with a flux limit  $f_{eff}$ . Note that the flux limit changes over the sky although the count rate limit is uniform, i.e.,  $0.1 \text{ cts s}^{-1}$ . This is because the flux limit depends sensitively on the column density. For the conversion of the count rate to flux, a spectrum of  $\Gamma = 2$  was assumed, while the appropriate  $N_H$  values were taken from Hartmann & Burton (1997). The effective RASSBSC flux limit is then the *average* flux limit over the areas covered by our sample; we find  $f_{eff} \sim 10^{-12} \text{ erg cm}^{-2} \text{ s}^{-1}$ .

We derive a value of  $r_0 = 6.7 \pm 1.0 h^{-1} \text{ Mpc}$  assuming comoving clustering evolution and freezing the parameter  $\gamma$  at the value of 1.8. Alternatively, if we repeat our calculations using  $\gamma \approx 1.9$ , which we derived from our unconstrained  $w(\theta)$  fit, we find  $r_0 = 5.4 \pm 0.9 h^{-1} \text{ Mpc}$ . We consider the difference  $\delta r_0 \approx 1.3 h^{-1} \text{ Mpc}$  between these two determinations of  $r_0$  as a further source of uncertainty and we quote as our best estimate of  $r_0$  their average value:  $r_0 \sim 6.0 \pm 1.6 h^{-1} \text{ Mpc}$ . The quoted uncertainty is derived by adding in quadrature  $\delta r_0$  and the formal fitting errors. Note that in the case of stable clustering and for  $\gamma = 1.8$  we obtain  $r_0 = 6.5 \pm 1.0 h^{-1} \text{ Mpc}$ .

#### 4 DISCUSSION AND CONCLUSIONS

Assuming a  $\gamma = 1.8$  power-law model we detect a significant clustering at  $4.1\sigma$  confidence level between 0 and 8 degrees with  $w(\theta < 8^\circ) = 2.5 \pm 0.6 \times 10^{-2}$ . Our analysis of  $w(\theta)$  implies  $r_0 = 6.0 \pm 1.6 h^{-1} \text{ Mpc}$ . Note that our result comes from scales larger than  $\sim 2$  degrees which corresponds to  $\sim 10 h^{-1} \text{ Mpc}$  at a redshift of  $z \sim 0.1$ . Although the statistics of our sample are limited, the lack of signal at small scales is significant at a high confidence level. This could imply a change in the form of the AGN correlation function at small separations. Our result represents the first significant detection of clustering in the local X-ray Universe. The Boyle & Mo (1993) analysis of a sample of 183 low-redshift ( $z < 0.2$ ) AGNs, from the EMSS survey, represented the first X-ray selected AGN clustering study in the local Universe. They detected a weak but not significant clustering signal on scales  $< 10^{-1} h^{-1} \text{ Mpc}$ .

Comparison of our findings with previous results at high redshift can give important clues on the evolution of clustering in the Universe. We first compare with the results of Carrera et al. (1998) who analysed the clustering properties of two *ROSAT*, soft X-ray selected AGN samples at higher redshifts, the *RIXOS* and the *ROSAT* Deep Survey samples containing about 235 AGN in total. They detect a  $2\sigma$  clustering signal in the *RIXOS* sample (mean  $z \sim 0.8$ ) while no clustering is detected in the DRS sample. A correlation length of  $< 3.5 h^{-1} \text{ Mpc}$  was found in the case of

comoving evolution, significantly lower than our results. However, as their derived correlation length,  $r_0$ , may depend upon  $r_c$  (the distance up to which the number of predicted and observed pairs are compared in order to derive  $r_0$ ), larger samples are needed at high redshifts in order to constrain  $r_0$ . Better statistics can be provided by using optical samples. Croom & Shanks (1996) used a sample of  $\sim 1700$  UVX AGN. They obtain a comoving clustering length of  $r_0 = 5.4 \pm 1.1 h^{-1}$  comoving Mpc at a mean redshift of  $z = 1.3$ . Therefore, comparison with our clustering length at low redshifts provides strong evidence for the comoving model of clustering evolution.

Our derived AGN clustering length is similar to that of local galaxies. Hence it appears that AGN randomly sample the galaxy distribution and they do not trace the high peaks of the density field despite their low density and high luminosity (cf. Efstathiou & Rees 1988). This result is corroborated by the analysis of the cross-correlation of low-redshift EMSS AGN with APM galaxies (Smith, Boyle & Maddox 1995). They find that the cross-correlation amplitude is similar to the galaxy auto-correlation amplitude thus supporting the idea that AGN randomly sample the galaxy population. This is again consistent with imaging studies of the AGN environments (cf. Boyle & Couch 1993). These authors find no excess number of galaxies around radio-quiet AGN, again pointing out that galaxies and AGN have similar environments.

Our work has important implications for the origin of the X-ray background (XRB). Several studies constrained the contribution of AGN to the XRB by using the auto-correlation function of the XRB fluctuations (e.g. Carrera & Barcons 1992; Georgantopoulos et al. 1993; Soltan & Hasinger 1994). These studies conclude that if AGN have a clustering length of  $6 h^{-1} \text{ Mpc}$  similar to optically selected AGN and they evolve according to the comoving model they can produce only about half of the XRB intensity. Our results appear to maintain the validity of this constraint.

In the near future, we anticipate a drastic improvement of our knowledge of AGN clustering, at least in X-ray wavelengths. The RASS identification programmes will be essential in providing a large AGN sample in the local Universe with redshift information. The new *ABRIXAS* mission will be invaluable in providing local AGN samples in the hard X-ray band which is not prone to Galactic absorption. Finally, the *XMM* serendipitous surveys are expected to provide tens of thousands high-redshift AGN thus strengthening the clustering statistics at high redshift.

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